## The Dirichlet-Discrete Model

Readings - Bishop: Section 2.2 and Appendix E The first six pages of the optional reading by Frigyik, Kapila, and Gupta are also recommended

- The third class discussed the Beta-Bernoulli Model
- This class will generalize that model from binary random variables to variables taking values in a finite set (often called "categorical" or "discrete" variables)
- For example, the set of words in a vocabulary
- For simplicity, we will denote this set as {1, 2, ..., V} where V is at least 2 and known in advance (the case where V is not known in advance is a topic for a more advanced class)

#### **Discrete Random Variables**

- The Beta-Bernoulli model had a single parameter,  $\boldsymbol{\mu}$
- Now  $\mu$  becomes a vector with V components:  $\mu = [\mu_1, \mu_2, ..., \mu_V]$  with  $\mu_i \ge 0$  and  $\sum \mu_i = 1$
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  [0, 0, ..., 0, 1, 0, ..., 0], where the position of the 1 indicates the value of the variable
- If  $X_i$  is a discrete random variable, we can express its probability distribution as  $DiscretePMF(X = w) = \prod_i \mu_i^{X_{wi}}$ where  $X_{wi} = 1$  only when  $X_i = w$

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- This gives the likelihood function

 $P(X_1, X_2, ..., X_N | \mu) = \prod_n \prod_i \mu_i^{X_{ni}} = \prod_i \mu_i^{m_i}$ 

where  $m_i = \sum_n X_{ni}$  is a count of the number of times word i appears in the dataset

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- $\mu^{ML} = \arg \max \sum_{i} m_{i} \ln \mu_{i}$ where the maximum is over all  $\mu$  in  $\Delta^{V}$
- Since Δ<sup>V</sup> is the V-1 dimensional subspace of legal V dimensional probability vectors, this is a constrained optimization problem and we can use *Lagrange multipliers* to find the µ that gives the maximum of the likelihood function

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- This motivates the definition of the Lagrangian:  $\mathcal{L}(\mathbf{x}, \lambda) = f(\mathbf{x}) + \lambda g(\mathbf{x})$
- To maximize  $f(\mathbf{x})$  subject to  $g(\mathbf{x}) = 0$ , take partial derivatives of  $\mathcal{L}(\mathbf{x}, \lambda)$  with respect to  $\lambda$  and the components of  $\mathbf{x}$  and set these derivatives to zero

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- Plugging these values of  $\mu_i$  into the constraint gives  $\lambda = \sum_i m_i = N$
- Putting this all together gives  $\mu^{ML} = [m_1/N, m_2/N, ..., m_V/N]$ which has all its components in the interval [0,1] as desired

Sufficient Statistics and the Multinomial Distribution

• This means that (under the bag-of-words assumption) all we need to know about the data is contained in the quantities  $m_i$  so the  $m_i$  are called *sufficient statistics* for  $\mu^{ML}$ 

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- The distribution of the  $m_i$  values, conditioned on  $\mu$  and N is *multinomial*:

Mult(m<sub>1</sub>, m<sub>2</sub>, ..., m<sub>V</sub> |  $\mu$ , N) = C(N; m<sub>1</sub>, m<sub>2</sub>, ..., m<sub>V</sub>)  $\prod_{i} \mu_{i}^{m_{i}}$ 

where  $C(N; m_1, m_2, ..., m_V) = N! / (m_1! m_2! ... m_V!)$ are the *multinomial coefficients* found in the expansion of  $(x_1 + x_2 + ... + x_V)^N$ 

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• Continuing the analogy to the Beta-Bernoulli model, we can generalize the multinomial distribution to the *Dirichlet* distribution, again replacing the factorials with gamma functions

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DirPDF( $\mu | a_1, a_2, ..., a_V$ ) = c(**a**)  $\prod_i \mu_i^{a_i - 1}$ 

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- When V=2, this is the Beta distribution
- The Dirichlet distribution is defined on the probability simplex given by the constraints  $\sum_i \mu_i = 1$  and  $\mu_i \ge 0$

• The Dirichlet-Discrete joint distribution defines a complete model:

 $P(X_1, X_2, ..., X_N, \mu) =$ 

[ $\prod_n$  DiscretePMF(X<sub>n</sub> |  $\mu$ )] DirPDF( $\mu$  |  $a_1$ ,  $a_2$ , ...,  $a_v$ ) where the first factor is the likelihood and the second is the prior

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- Predictive Posterior:  $P(X_N | X_1, X_2, ..., X_{N-1}) = \int_{\Delta V} P(X_N | \boldsymbol{\mu}) P(\boldsymbol{\mu} | X_1, X_2, ..., X_{N-1}) d\boldsymbol{\mu}$ 

Suppose we have some prior knowledge of µ, represented as a prior distribution, and we want to combine this with new data X<sub>1</sub>, X<sub>2</sub>, ..., X<sub>N</sub>, to obtain a posterior distribution for µ and use this to get a *maximum a posteriori* (MAP) estimate for µ

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- The evidence, P(X<sub>1</sub>, X<sub>2</sub>, ..., X<sub>N</sub>), and the normalizing factor, c(a), don't depend on μ, so they don't affect the MAP estimate:

 $\begin{array}{l} P(\boldsymbol{\mu} \mid X_1, X_2, \ldots, X_N) \sim \\ [\prod_n \text{DiscretePMF}(X_n \mid \boldsymbol{\mu})] \text{ DirPDF}(\boldsymbol{\mu} \mid a_1, a_2, \ldots, a_V) \sim \end{array}$ 

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where  $\sim$  means that factors not involving  $\mu$  have been omitted

- $\mu^{MAP} = \arg \max P(\mu \mid X_1, X_2, ..., X_N)$ 
  - $= \arg \max \prod_{i} \mu_{i}^{m_{i}+a_{i}-1}$
  - = arg max  $\sum_{i}(m_i+a_i-1) \ln \mu_i$

•  $\boldsymbol{\mu}^{\text{MAP}} = \arg \max P(\boldsymbol{\mu} \mid X_1, X_2, ..., X_N)$ =  $\arg \max \prod_i \mu_i^{m_i + a_i - 1}$ 

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• Note that the  $a_i$  need to be at least 1 to ensure that the coefficients ( $m_i+a_i-1$ ) are non-negative

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- Given that

 $\mu^{MAP} = \arg \max \sum_{i} (m_i + a_i - 1) \ln \mu_i$ 

we can again use Lagrange multipliers to obtain

 $\boldsymbol{\mu}^{\text{MAP}} = [(m_1 + a_1 - 1)/(N + \sum_i (a_i - 1), \dots, (m_v + a_v - 1)/(N + \sum_i (a_i - 1))]$