## The Dirichlet-Discrete Model

Readings - Bishop: Section 2.2 and Appendix E The first six pages of the optional reading by Frigyik, Kapila, and Gupta are also recommended

- The third class discussed the Beta-Bernoulli Model
- This class will generalize that model from binary random variables to variables taking values in a finite set (often called "categorical" or "discrete" variables)
- For example, the set of words in a vocabulary
- For simplicity, we will denote this set as $\{1,2, \ldots, \mathrm{~V}\}$ where V is at least 2 and known in advance (the case where V is not known in advance is a topic for a more advanced class)


## Discrete Random Variables

- The Beta-Bernoulli model had a single parameter, $\mu$
- Now $\mu$ becomes a vector with V components:

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\boldsymbol{\mu}=\left[\mu_{1}, \mu_{2}, \ldots, \mu_{\mathrm{v}}\right] \text { with } \mu_{\mathrm{i}} \geq 0 \text { and } \sum \mu_{\mathrm{i}}=1
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- The value of a discrete random variable can be represented by a "one-hot" vector with V components: $[0,0, \ldots, 0,1,0, \ldots, 0]$, where the position of the 1 indicates the value of the variable
- If $X_{i}$ is a discrete random variable, we can express its probability distribution as $\operatorname{DiscretePMF}(\mathrm{X}=\mathrm{w})=\prod_{\mathrm{i}} \mu_{\mathrm{i}}{ }^{\mathrm{X}_{\mathrm{wi}}}$ where $\mathrm{X}_{\mathrm{wi}}=1$ only when $\mathrm{X}_{\mathrm{i}}=\mathrm{w}$


## The Likelihood Function

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- This gives the likelihood function

$$
\mathrm{P}\left(\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}} \mid \boldsymbol{\mu}\right)=\prod_{\mathrm{n}} \prod_{\mathrm{i}} \mu_{\mathrm{i}}^{\mathrm{X}_{\mathrm{ni}}}=\prod_{\mathrm{i}} \mu_{\mathrm{i}}^{\mathrm{m}_{\mathrm{i}}}
$$

where $m_{i}=\sum_{\mathrm{n}} \mathrm{X}_{\mathrm{ni}}$ is a count of the number of times word i appears in the dataset

## A Maximum Likelihood Estimate for $\boldsymbol{\mu}$

- Since $\ln (x)$, the natural logarithm of $x$, is an increasing function of $x$, we can maximize the log-likelihood instead of maximizing the likelihood directly. This simplifies the math and helps prevent numerical problems.


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where the maximum is over all $\boldsymbol{\mu}$ in $\Delta^{\mathrm{V}}$
- Since $\Delta^{\mathrm{V}}$ is the $\mathrm{V}-1$ dimensional subspace of legal V dimensional probability vectors, this is a constrained optimization problem and we can use Lagrange multipliers to find the $\boldsymbol{\mu}$ that gives the maximum of the likelihood function


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- To maximize $f(\mathbf{x})$ subject to $g(\mathbf{x})=0$, take partial derivatives of $\mathcal{L}(\mathbf{x}, \lambda)$ with respect to $\lambda$ and the components of $\mathbf{x}$ and set these derivatives to zero

Using Lagrange Multipliers to Find a Maximum Likelihood Estimate for $\boldsymbol{\mu}$

- First write the Lagrangian:

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\mathcal{L}\left(\mu_{1}, \mu_{2}, \ldots, \mu_{\mathrm{v}}, \lambda\right)=\sum_{\mathrm{i}} \mathrm{~m}_{\mathrm{i}} \ln \mu_{\mathrm{i}}+\lambda\left(1-\sum_{\mathrm{i}} \mu_{\mathrm{i}}\right)
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- Putting this all together gives $\boldsymbol{\mu}^{\mathrm{ML}}=\left[\mathrm{m}_{1} / \mathrm{N}, \mathrm{m}_{2} / \mathrm{N}, \ldots, \mathrm{m}_{\mathrm{V}} / \mathrm{N}\right]$ which has all its components in the interval $[0,1]$ as desired


## Sufficient Statistics and the Multinomial Distribution

- This means that (under the bag-of-words assumption) all we need to know about the data is contained in the quantities $\mathrm{m}_{\mathrm{i}}$ so the $\mathrm{m}_{\mathrm{i}}$ are called sufficient statistics for $\boldsymbol{\mu}^{\mathrm{ML}}$


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- The distribution of the $\mathrm{m}_{\mathrm{i}}$ values, conditioned on $\boldsymbol{\mu}$ and N is multinomial:

$$
\operatorname{Mult}\left(\mathrm{m}_{1}, \mathrm{~m}_{2}, \ldots, \mathrm{~m}_{\mathrm{v}} \mid \boldsymbol{\mu}, \mathrm{N}\right)=\mathrm{C}\left(\mathrm{~N} ; \mathrm{m}_{1}, \mathrm{~m}_{2}, \ldots, \mathrm{~m}_{\mathrm{v}}\right) \prod_{\mathrm{i}} \mu_{\mathrm{i}^{\mathrm{m}}}
$$

where $\mathrm{C}\left(\mathrm{N} ; \mathrm{m}_{1}, \mathrm{~m}_{2}, \ldots, \mathrm{~m}_{\mathrm{v}}\right)=\mathrm{N}!/\left(\mathrm{m}_{1}!\mathrm{m}_{2}!\ldots \mathrm{m}_{\mathrm{v}}!\right)$ are the multinomial coefficients found in the expansion of

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\left(\mathrm{x}_{1}+\mathrm{x}_{2}+\ldots+\mathrm{xv}_{\mathrm{v}}\right)^{\mathrm{N}}
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\left(x_{1}+x_{2}+\ldots+x_{v}\right)^{N}
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- Continuing the analogy to the Beta-Bernoulli model, we can generalize the multinomial distribution to the Dirichlet distribution, again replacing the factorials with gamma functions


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- When $\mathrm{V}=2$, this is the Beta distribution
- The Dirichlet distribution is defined on the probability simplex given by the constraints $\sum_{i} \mu_{i}=1$ and $\mu_{i} \geq 0$


## Distributions Derived from the Dirichlet-Discrete Model

- The Dirichlet-Discrete joint distribution defines a complete model:
$\mathrm{P}\left(\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}, \boldsymbol{\mu}\right)=$
$\left[\prod_{n} \operatorname{DiscretePMF}\left(X_{n} \mid \boldsymbol{\mu}\right)\right] \operatorname{DirPDF}\left(\boldsymbol{\mu} \mid a_{1}, a_{2}, \ldots, a_{v}\right)$
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- Several distributions can be derived from this:
— Evidence: $\mathrm{P}\left(\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}\right)=\int_{\Delta \mathrm{V}} \mathrm{P}\left(\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}, \boldsymbol{\mu}\right) \mathrm{d} \boldsymbol{\mu}$


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- Posterior: $\mathrm{P}\left(\boldsymbol{\mu} \mid \mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}\right)$, obtained by dividing the joint distribution by the evidence (Bayes rule)


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- Predictive Posterior: $\mathrm{P}\left(\mathrm{X}_{\mathrm{N}} \mid \mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}-1}\right)=$

$$
\int_{\Delta V} P\left(X_{N} \mid \boldsymbol{\mu}\right) P\left(\boldsymbol{\mu} \mid X_{1}, X_{2}, \ldots, X_{N-1}\right) d \boldsymbol{\mu}
$$

## Deriving the MAP Estimate from the Posterior

- Suppose we have some prior knowledge of $\boldsymbol{\mu}$, represented as a prior distribution, and we want to combine this with new data $\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}$, to obtain a posterior distribution for $\boldsymbol{\mu}$ and use this to get a maximum a posteriori (MAP) estimate for $\boldsymbol{\mu}$


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- The evidence, $\mathrm{P}\left(\mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}\right)$, and the normalizing factor, $\mathrm{c}(\mathbf{a})$, don't depend on $\boldsymbol{\mu}$, so they don't affect the MAP estimate:
$\mathrm{P}\left(\boldsymbol{\mu} \mid \mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}\right) \sim$
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- $\boldsymbol{\mu}^{\mathrm{MAP}}=\arg \max \mathrm{P}\left(\boldsymbol{\mu} \mid \mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{N}}\right)$
$=\arg \max \prod_{i} \mu_{i^{m i}+\mathrm{a}_{\mathrm{i}}-1}$ $=\arg \max \sum_{i}\left(\mathrm{~m}_{\mathrm{i}}+\mathrm{a}_{\mathrm{i}}-1\right) \ln \mu_{\mathrm{i}}$


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- Note that the $a_{i}$ need to be at least 1 to ensure that the coefficients $\left(m_{i}+a_{i}-1\right)$ are non-negative
- Given that

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\boldsymbol{\mu}^{\mathrm{MAP}}=\arg \max \sum_{\mathrm{i}}\left(\mathrm{~m}_{\mathrm{i}}+\mathrm{a}_{\mathrm{i}}-1\right) \ln \mu_{\mathrm{i}}
$$

we can again use Lagrange multipliers to obtain

$$
\boldsymbol{\mu}^{\mathrm{MAP}}=\left[\left(\mathrm{m}_{1}+\mathrm{a}_{1}-1\right) /\left(\mathrm{N}+\sum_{\mathrm{i}}\left(\mathrm{a}_{\mathrm{i}}-1\right), \ldots,\left(\mathrm{m}_{\mathrm{v}}+\mathrm{a}_{\mathrm{v}}-1\right) /\left(\mathrm{N}+\sum_{\mathrm{i}}\left(\mathrm{a}_{\mathrm{i}}-1\right)\right]\right.\right.
$$

